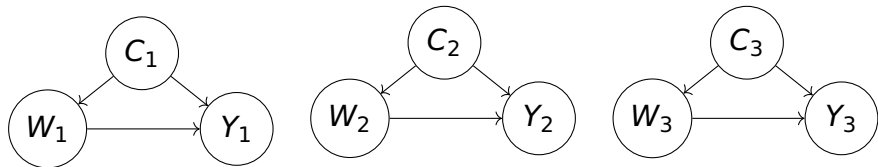


Treatment Effect Estimation with Interference and Confounding using Graphical Models

Leonard Henckel, Meta-Lina Spohn and Marloes H. Maathuis

Stable unit treatment value assumption (SUTVA):
(Potential) outcome on one unit should be unaffected by treatment assignments of other units



- No **interference**: W_i does not affect Y_j for $i \neq j$
- No contagion: Y_i does not affect Y_j for $i \neq j$

Examples with [interference](#):

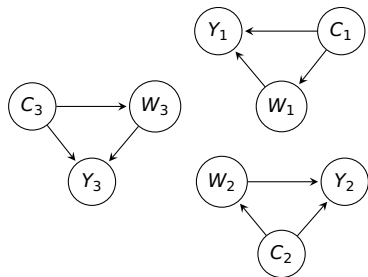
- Crosscontamination in agricultural studies
- Instruction type in school
- Vaccinations
- Communication of advertisement

Remark: Common were units can communicate/compete

Interference

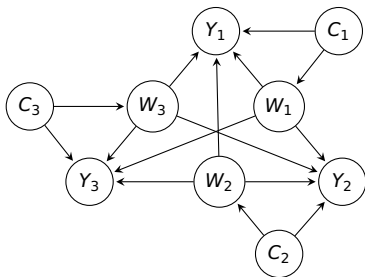
No interference

- $C_i \leftarrow \epsilon_{C_i}$
- $W_i \leftarrow g(C_i, \epsilon_{W_i})$
- $Y_i \leftarrow h(W_i, C_i, \epsilon_i)$



Interference

- $C_i \leftarrow \epsilon_{C_i}$
 - $W_i \leftarrow g(C_i, \epsilon_{W_i})$
 - $Y_i \leftarrow h(\mathbf{W}, C_i, \epsilon_i)$,
- $\mathbf{W} := (W_1, \dots, W_N)$.



Classical i.i.d. setting:

- Two possible treatment assignments
- Average Treatment Effect:
 $E[Y|do(W = 1)] - E[Y|do(W = 0)]$

Interference setting:

- 2^N possible treatment assignments
- Global Average Treatment Effect:
 $\tau = \frac{1}{N} \sum_{i=1}^N (E[Y_i|do(\mathbf{W} = \mathbf{1})] - E[Y_i|do(\mathbf{W} = \mathbf{0})])$
- Stochastic Interventions:
 $\tau(P_\pi, P_\eta) = \frac{1}{N} \sum_{i=1}^N (E[Y_i|do(\mathbf{W} \sim P_\pi)] - E[Y_i|do(\mathbf{W} \sim P_\eta)])$

Conceptual settings:

- Interference as nuisance to correct for
- [Interference as part of effect](#)

Approaches to handling interference:

- Experimental design (e.g. VanderWeele and Tchetgen Tchetgen, 2011)
- Partial interference: Arbitrary interference but within non-overlapping groups (e.g. Hudgens and Halloran, 2008)
- [Explicit modelling of interference](#) (e.g. Chin, 2019)

Remark: Without restrictions on the interference, treatment effects are generally not identifiable.

Modelling interference with features

Network graph G :

Graph describing interactions between units, e.g.,

- Friendship ties in a social group
- Geographical adjacency

Features: $\mathbf{X} = f(\mathbf{W}, G)$ (Manski, 1993; Chin, 2019)

Deterministic functions of G and treatment vector \mathbf{W} , e.g.,

- Fraction of treated neighbors
- Fraction of treated neighbors of neighbors

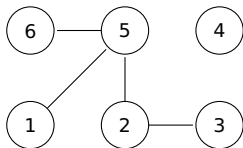
Modelling interference:

$$Y_i \leftarrow h(\mathbf{W}, \mathbf{C}_i, \epsilon_i) = h'(W_i, \mathbf{X}_i, \mathbf{C}_i, \epsilon_i)$$

Modelling interference with features

Example: exposure to advertisement

- $X_i = f(\mathbf{W}, G) =$ number of treated neighbors of i in G ,
- $Y_i \leftarrow \mu_i + \mathbf{W}_i + X_i + \epsilon_i$.



$$\begin{aligned}\tau &= \frac{1}{6} \sum_{i=1}^6 (E[Y_i | do(\mathbf{W} = \mathbf{1})] - E[Y_i | do(\mathbf{W} = \mathbf{0})]) \\ &= \frac{1}{6} \sum_{i=1}^6 (1 + E[X_i | do(\mathbf{W} = \mathbf{1})] - 0 - E[X_i | do(\mathbf{W} = \mathbf{0})]),\end{aligned}$$

where $E[X_i | do(\mathbf{W} = \mathbf{1})] = (1, 2, 1, 0, 3, 1)$.

$$\implies \tau = (2 + 3 + 2 + 1 + 4 + 2)/6 - 0 = 14/6.$$

Definition

Unit SEM with interference features & linear outcome model (inspired by Chin, 2019):

- $\mathbf{C}_i \leftarrow \mathcal{A}\mathbf{C}_i + \epsilon_{\mathbf{C}_i}$,
- $W_i \leftarrow g(\mathbf{C}_i, \epsilon_{W_i})$,
- $\mathbf{X}_i \leftarrow f(\mathbf{W}, G)$,
- $Y_i \leftarrow W_i\beta^{(1)}(\mathbf{1}, \mathbf{X}_i) + (1 - W_i)\beta^{(0)}(\mathbf{1}, \mathbf{X}_i) + \gamma\mathbf{C}_i + \epsilon_i$.

Remark: Outcome model corresponds to

$$Y_i \leftarrow \alpha^{(0)}(\mathbf{1}, \mathbf{X}_i) + \alpha^{(1)}(W_i, \mathbf{O}_i) + \gamma\mathbf{C}_i + \epsilon_i,$$

with $\beta^{(0)} = \alpha^{(0)}$, $\beta^{(1)} = \alpha^{(0)} + \alpha^{(1)}$ and $\mathbf{O}_i = W_i\mathbf{X}_i$.

Results:

- Reformulate GATE τ and stochastic effects $\tau(P_\pi, P_\eta)$
- Define new notion of Interference-DAG (I-DAG)
- Propose adjustment estimator based on graphical criteria on I-DAG
- Propose instrumental variable estimator based on graphical criteria on I-DAG
- Derive sufficient conditions for consistency
- Derive sufficient conditions for asymptotic normality

Lemma: (Stochastic treatment contrast with features)

$\tau_N(P_\pi, P_\eta) = \omega_N^{(1)}(\pi, \eta)(\alpha^{(0)} + \alpha^{(1)}) - \omega_N^{(0)}(\pi, \eta)\alpha^{(0)}$ where

$$\omega_N^{(1)} = \frac{1}{N} \sum_{i=1}^N (\pi E[X_i | do(W \sim P_\pi)] - \eta E[X_i | do(W \sim P_\eta)]), \text{ and}$$

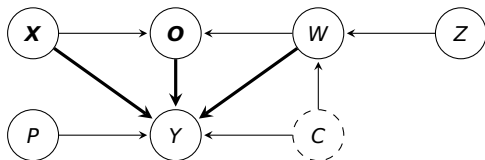
$$\omega_N^{(0)} = \frac{1}{N} \sum_{i=1}^N ((1 - \pi)E[X_i | do(W \sim P_\pi)] - (1 - \eta)E[X_i | do(W \sim P_\eta)])$$

Remark: Weights can be computed or estimated with Monte Carlo so only need to estimate $\alpha^{(0)}$ and $\alpha^{(1)}$.

Idea: interpret feature model as non-i.i.d. data from DAG

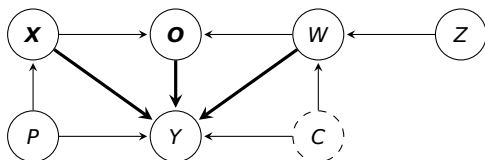
Example:

- $Z_i \leftarrow \epsilon_{Z_i}, P_i \leftarrow \epsilon_{P_i}, C_i \leftarrow \epsilon_{C_i},$
- $W_i \leftarrow g(Z_i, \epsilon_{W_i}),$
- $\mathbf{X}_i \leftarrow f(\mathbf{W}_{-i}, G),$
- $\mathbf{O}_i = W_i \mathbf{X}_i,$
- $Y_i = \alpha_0^{(0)} + \alpha_{1:P}^{(0)} \mathbf{X}_i + \alpha_0^{(1)} W_i + \alpha_{1:P}^{(1)} \mathbf{O}_i + \gamma_C \mathbf{C}_i + \gamma_P P_i + \epsilon_i.$



Example: Unit SEM:

- $Z_i \leftarrow \epsilon_{Z_i}, P_i \leftarrow \epsilon_{P_i}, C_i \leftarrow \epsilon_{C_i},$
- $W_i \leftarrow g(Z_i, \mathbf{P}_{-i}, \epsilon_{W_i}),$
- $\mathbf{X}_i \leftarrow f(\mathbf{W}_{-i}, G),$
- $\mathbf{O}_i = W_i \mathbf{X}_i,$
- $Y_i = \alpha_0^{(0)} + \alpha_{1:P}^{(0)} \mathbf{X}_i + \alpha_0^{(1)} W_i + \alpha_{1:P}^{(1)} \mathbf{O}_i + \gamma_C C_i + \gamma_P P_i + \epsilon_i.$



Definition (Interference-DAG (I-DAG))

- If $A_i \rightarrow B_i$ for some i in the unit graph, then $A \rightarrow B$ in the I-DAG.
- If $A_i \rightarrow W_j$ for some $i \neq j$ in the unit graph, then $A \rightarrow X$ in the I-DAG.

Lemma (I-DAG encodes independence statements)

If $\mathbf{A} \perp_{\mathcal{G}} \mathbf{B} \mid \mathbf{C}$ in the I-DAG, then $\mathbf{A}_i \perp\!\!\!\perp \mathbf{B}_i \mid \mathbf{C}_i$ for all $i = 1, \dots, N$.

Idea:

- $\alpha^{(0)}$ and $\alpha^{(1)}$ suffice to estimate $\tau_N(P_\pi, P_\eta) = \omega_N^{(1)}(\pi, \eta)(\alpha^{(0)} + \alpha^{(1)}) - \omega_N^{(0)}(\pi, \eta)\alpha^{(0)}$.
- $(\alpha^{(0)}, \alpha^{(1)})$ is total effect of pseudo-intervention on $(W, \mathbf{X}, \mathbf{O})$ on Y in I-DAG \mathcal{G} .
 \implies Apply graphical criteria from literature to I-DAG to estimate $\alpha^{(0)}$ and $\alpha^{(1)}$

Lemma

$\mathbf{P} = \text{pa}(Y, \mathcal{G}) \setminus \mathbf{T}$ and \mathbf{Z} valid adjustment set. Then $\beta_{yt.\mathbf{p}} = \beta_{yt.\mathbf{z}}$ where $\mathbf{P} = \text{pa}(Y, \mathcal{G}) \setminus \{X\}$ and $\mathbf{T} = \{W\} \cup \mathbf{X} \cup \mathbf{O}$.

Assumption 1:

Assume existence of:

- $\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N E[X_i | do(W \sim P_\pi)]$
- $\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N E[X_i | do(W \sim P_\eta)]$.

Assumption 2:

Constrain dependencies between units by using yet another auxiliary graph called [dependency graph](#).

Remark: Implicitly constrains underlying network graph without explicitly modelling it.

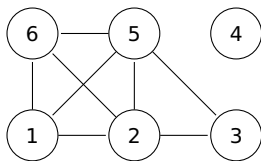
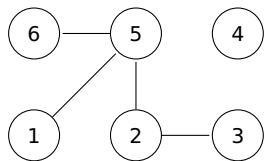
Dependency graph

Definition: (Interference dependency graph)

$U_i \rightarrow U_j$ present in $D(\mathbf{X}, \mathbf{W})$ iff:

- W_i affects \mathbf{X}_j ,
- \mathbf{X}_i and \mathbf{X}_j are affected by some W_k , $k \neq i, j$.

Example: X_i = number of treated neighbors of i in G .



Adjustment consistency

Theorem: (Consistency for adjustment estimator)

\mathbf{D} valid adjustment set relative to (\mathbf{A}, Y) with $\mathbf{A} = (\mathbf{X}, W, \mathbf{O})$ in I-DAG \mathcal{G} , $\tilde{\mathbf{X}} = (\mathbf{A}, \mathbf{D})^T$. If

- $E[Y_i^4] < \infty$ and $E[\tilde{\mathbf{X}}_i^4] < \infty \forall i = 1, \dots, N$
- $E[\tilde{\mathbf{X}}_i \tilde{\mathbf{X}}_i^T] < \infty$ is invertible $\forall i = 1, \dots, N$,
- $\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N E[\tilde{\mathbf{X}}_i \tilde{\mathbf{X}}_i^T] = \Sigma_{\tilde{\mathbf{X}}\tilde{\mathbf{X}}} < \infty$, where $\Sigma_{\tilde{\mathbf{X}}\tilde{\mathbf{X}}}$ is invertible,
- $d(N) \in o(N)$, where $d(N)$ is the maximal degree in the dependency graph $D(X, W)$,

then

$$\tau_{ols}(P_\pi, P_\eta) - \tau_N(P_\pi, P_\eta) \xrightarrow{P} 0,$$

Remark: $d(N) \in o(N) \implies$ no non-local features

Adjustment asymptotic normality

Theorem: (Asymptotic normality of adjustment estimator)
 \mathbf{D} valid adjustment set relative to (\mathbf{A}, Y) with $\mathbf{A} = (\mathbf{X}, W, \mathbf{O})$ in I-DAG \mathcal{G} , $\tilde{\mathbf{X}} = (\mathbf{A}, \mathbf{D})^T$. If

- $E[Y_i^8] < \infty$ and $E[\tilde{\mathbf{X}}_i^8] < \infty \forall i = 1, \dots, N$,
- $E[\tilde{\mathbf{X}}_i \tilde{\mathbf{X}}_i^T] < \infty$ is invertible $\forall i = 1, \dots, N$,
- $\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N E[\tilde{\mathbf{X}}_i \tilde{\mathbf{X}}_i^T] = \Sigma_{\tilde{\mathbf{X}}\tilde{\mathbf{X}}} < \infty$, where $\Sigma_{\tilde{\mathbf{X}}\tilde{\mathbf{X}}}$ is invertible,
- $\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N E[\epsilon_i^2 \tilde{\mathbf{X}}_i \tilde{\mathbf{X}}_i^T] = \Sigma_{\epsilon^2 \tilde{\mathbf{X}}\tilde{\mathbf{X}}} < \infty$, where
 $\epsilon_i := Y_i - \tilde{\mathbf{X}}_i^T (\gamma_A, \gamma_D)^T$,
- $d(N) \in o(N^{1/4})$,

then

$$\sqrt{N} (\tau_{OLS}(P_\pi, P_\eta) - \tau_N(P_\pi, P_\eta)) \xrightarrow{D} \mathcal{N}(0, \sigma^2),$$

Summary

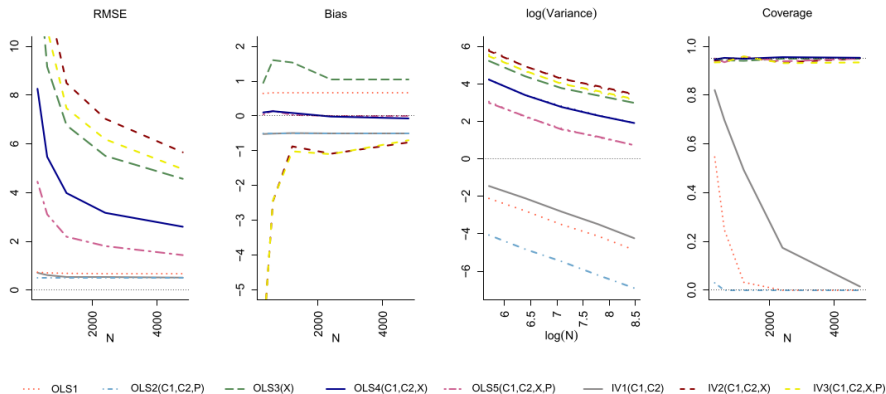
- Framework that allows for both confounding and interference
- I-DAG to make classical graphical results usable
- Adjustment estimator
- Instrumental variables estimator
- Necessary conditions for consistency and asymptotic normality

Thanks!

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Simulations



\mathcal{G} : Watts-Strogatz network, Target: $\tau(P_\pi, P_\eta)$ with $\pi = 0.7$ and $\eta = 0.2$, \mathbf{X} : i) fraction of treated neighbors and ii) fraction of treated neighbors of neighbors.